# The Validity of Ear Prediction Guidelines Used in Facial Approximation*,†,* 


#### Abstract

This study examined eight previously published ear prediction methods by Welcker, Gerasimov, Fedosyutkin and Nainys, and Broadbent and Mathews. Computed tomography scans of 78 living adults $\left(n_{1}\right)$ did not support any of these previously published recommendations. Free earlobes were found to accompany protruding supramastoid crests (Pearson's $\chi^{2}<0.05$ ); and ear length [l] and width [w] differed by sex ( $p<0.05$ ), correlated with age ( $r=0.38[1] ; 0.32[\mathrm{w}]$ ), and correlated with facial height $(r=0.37[1] ; 0.30[\mathrm{w}]$ ). New regression equations (for ear length and width) were generated using these variables in several samples and, where possible, cross-validated using independent data ( $n_{1}=78, n_{2}=2190$, $n_{3}=1328, n_{4}=1010$, and $n_{5}=47$ ). As a result of these analyses, four valid and tested methods for ear prediction were identified, but large degrees of error continue to make accurate prediction of the ear, from the skull, problematic.


KEYWORDS: forensic science, facial reconstruction, facial reproduction, external ear, pinna, skull

Facial approximation is used to estimate a face from a bare skull under blind conditions, that is, without a priori knowledge of the individual's facial appearance. Irrespective of the specific method used, the long-standing aim of facial approximation has been to produce a face, from the skull, that is correctly and easily recognized (1-9). Not only does this goal give facial approximation its greatest utility for the resolution of forensic cases, but as a duality it also necessitates the use of accurate soft-tissue prediction methods.

Over the last 10 years, tests have revealed that established face prediction methods perform suboptimally. This includes large inaccuracies for popular methods to determine mouth width $(10,11)$, eyeball position (12-15), nose projection $(16,17)$, and the

[^0]temporalis muscle form (18). This has led to improved methods to determine these facial components (see the aforementioned studies and also [19-21]). However, the accuracy of preexisting and untested methods for other facial regions, including the ears, is unknown.

As part of the ongoing work to improve facial approximation methods, these other guidelines must be verified. Even if ears may contribute less to facial recognition than other facial components such as the eyes $(22,23)$, evaluation of these prediction guidelines is important as auricles nevertheless are required for overall correct gestalt appearance of the face. Moreover, their importance is probably elevated for persons possessing more atypical features (e.g., large protruding ears). Although multiple ear prediction guidelines exist, few of these have been subject to empirical validation.
The assessment of the ear in relation to the skull, at least as it pertains to facial approximation and craniofacial superimposition work, was first undertaken by Welcker (24). Welcker (24) stated that the cartilaginous opening of the ear is placed more posterior and superior than the bony opening (mean displacement $=5.3 \mathrm{~mm}$ )—a guideline that has stood up to scrutiny (25). In addition, Welcker (24) was the first to state that the main axis of the ear is parallel to the ascending ramus of the mandible (Fig. 1a)—a guideline that has been extensively cited in the published facial approximation literature, but without empirical backing from experimental data (see, e.g., [5,9,26-28]).

In 1955, Gerasimov added that the ear length (=height) roughly approximates the height of the nose measured from the base to the glabella (Fig. 1b) -a recommendation based, in part, on data collected from 462 Tajiks by Ginzburg (9). In later years, Gerasimov modified this rule to the ear representing the height of the nose from glabella to subnasale, plus an additional 2 mm (29). By using the distance between nasion to nasospinale on the skull, Jordanov (28) suggests that Gerasimov's directions can be directly applied to skulls, and concludes from measurements of 161 Bulgarians that


FIG. 1-Illustrations depicting the proposed equality between the angle of the jaw line and the main axis of the ear (a), and between the height of the ear and the length of the nose (b). Figures extracted from Gerasimov (9).

Gerasimov's rules are correct: that is, only slight differences existed between the "height of the nose" and the height of the ear (c. 2 mm ).

Other practitioners have also suggested that the height of the nose equals the height of the ear without the glabella qualifier (30-32), implying that the height of the nose represents the distance from sellion to subnasale. This latter guideline cannot be valid, however, because $95 \%$ of people possess ears that are, on average, c. 10 mm larger than their noses $(33,34)$. Krogman and Iscan (31) have also stated that the ear's height should equal 50 mm ; however, numerous metric studies contradict this suggestion, with mean ear height measurements in the vicinity of $60-70 \mathrm{~mm}$ (see, e.g., data summarized by Martin [35, p. 572], or Table 1 for more contemporary publications).

Other guidelines that Gerasimov (9) suggested, many of which have been restated by Fedosyutkin and Nainys (32) and Jordanov (28), include the following: (i) small and medially directed mastoid processes reflect small ears close to the head; (ii) massive and prominent mastoid processes denote a large and spread ears; (iii) a strongly developed supramastoid crest indicates a laterally projecting superior half of the ear; (iv) a rough external surface of the mastoid process indicates a laterally projecting inferior half of the ear; (v) a combination of the skeletal characteristics in "iv" and " v " gives a completely protruding ear (i.e., at both upper and lower poles); and (vi) the breadth of the ear equals half of its length. Note that this latter guideline appears to be inaccurate since typical ear index values reported in the early anthropological literature c. 0.60 (for summary, see Martin [35, p. 573]).

Broadbent and Mathews (27) in a review of artistic guidelines pertinent to surgical plastic facial reconstruction (that have been adopted for facial approximation) stated that the main axis of the ear was parallel to the angle of the dorsum of the soft tissue nose. This rule has since been invalidated by both Skiles and Randall (47) and Farkas et al. $(33,34,48)$ who found that the ear inclination is as much as $15^{\circ}$ more toward the vertical.

Fedosyutkin and Nainys (32) have also added that (i) the earlobe is free if the mastoid process is pointing forward (skull in Frankfurt

Horizontal) and (ii) the earlobe is attached if the mastoid process points more inferiorly. These observations were reportedly based on more than 200 skulls of identified individuals for whom facial photographs existed; however, trait frequencies and actual measurements are not provided in the original article (32).

It is worth noting here that the early anthropometric literature identifies differences in ear measurements according to sex (males larger than females) and the size of earlobe between individuals (for summary, see Martin [35]). The common, but typically informal, observation that ear size increases with age, and is associated with body size, has also been subject to recent empirical studies (40,44,45); however, these investigations have typically gone unmentioned in the facial approximation literature. Regression equations for ear length prediction have also been reported (see Table 2), but are limited because their errors are rarely fully disclosed, cross-validation results have not been pursued, and the practical value of the regressions, as relevant to facial approximation, has not been considered.

Recently, two further regression equations have been proposed to predict ear dimensions for males and females based on skeletal facial height (see equations no. 5 and 6 in Table 2, after Balueva et al. [49]). However, indicators of the accuracy of these equations (e.g., confidence intervals, $r^{2}$, or standard errors of the estimate [SEE]) are not presented by the original authors (49). This makes it difficult to judge the value of the methods, especially because these equations were formulated on the basis of correlations to facial height measurements on living subjects, not facial heights on skulls as the regression equations require.
This article aims to clarify the validity of some of the abovementioned ear prediction methods and attempts improvements using recently acquired medical computed tomography (CT) scans of living adults. To provide comprehensive assessments, four other samples of ear data are also drawn upon for cross-validation tests of regression equations. These data sets include (i) 2190 threedimensional (3D) facial scans of U.K. individuals by Evison and Vorder Bruegge (50); (ii) caliper measurements of 1328 White

TABLE 1—Mean adult ear heights ( mm ) published in the literature.
Where sides were not combined in the original study, the information for the left side has been reported.

| Mean | SD | Min | Max | $n$ | Age (Years) |  | Sample | Study |
| :---: | :--- | :--- | :--- | :--- | :--- | :--- | :--- | :--- |

*Weighted mean and standard deviation calculated by combining the original samples of 31-40, 41-50, 51-64, and 65-80 years.
${ }^{\dagger}$ Weighted mean calculated by combining original samples of $30-40,40-50,50-60$, and $60-70$ years.

American cadavers by T. W. Todd (unpublished data, courtesy of the Cleveland Museum of Natural History); (iii) caliper measurements of 1010 Black American cadavers by T. W. Todd (unpublished data, courtesy of the Cleveland Museum of Natural History); and (iv) caliper measurements of 47 contemporary living subjects taken specifically for this study.

## Material and Methods

CT scans of 78 living subjects $\left(n_{1}\right), 43$ males and 35 females, of known age (mean age $=41$ years; $\mathrm{SD}=19$ years; range $=18-$ 84 years) were used in this study so that soft tissue profiles of the ears, and their relationship to the skulls of the same individuals, could be examined. The scans represented medical images, collected from French Hospitals, with the subjects in the supine position and without any mechanical pressure on the ears. Patients presenting trauma or pathologies that affected the facial region were excluded. This sample, thereby, represents a much smaller component of a larger cohort, that is, CT scans of 500 individuals that were all screened for suitability to be used in this study. Ethical approval for the collection of the CT data was obtained from the French Comité de protection des personnes (Sud-Ouest et Outre Mer III). The slice thickness of scans in this sample was variable, but commonly $<1 \mathrm{~mm}$ (range $=0.6-1.4 \mathrm{~mm}$ ).

CT scans were viewed using the Treatment and Increased Vision for Medical Imaging (TIVMI) software developed by Bruno Dutailly (UMR 5199 PACEA). This software is freely downloadable (http://www.pacea.u-bordeaux1.fr/TIVMI/) and in addition to possessing classic geometric operators (such as planes, lines, segments, and outlines), it holds the advantage that 3D surface reconstruction can be undertaken using the Half-Maximum

Height protocol (51). This allows for a more accurate detection of the interface between tissues such as bone and soft tissue in contrast to other methods such as Marching Cubes algorithm (52).

Once the osseous and cutaneous surfaces were rendered from the DICOM files in the TIVMI software, landmarks were established in 3D both on the soft tissues of the ear and on the skull (Tables 3 and 4, Figs 2-4). In addition, morphoscopic assessment of the size of the supramastoid crest (Fig. 5) was used to evaluate the covariation of this feature with lateral ear protrusion. The morphology of the earlobe (attached or free) was also morphoscopically determined (Fig. 6) to evaluate its covariation with the skull following the recommendations of Fedosyutkin and Nainys (32).

Table 5 and Figs 2-4 present the measurements (angles and linear distances) that were used to evaluate the following preestablished ear prediction rules:
(i) The main axis of the ear is parallel to the ascending ramus of the mandible (24), that is, EA compared to MRA.
(ii) The height of the ear approximates the height of the nose ( $9,30-32$ ), that is, sa-sba compared to n -ss, $\mathrm{g}^{\prime}$-sn, se-sn, etc.
(iii) The height of the ear approximates the height of the nose from glabella to subnasale plus 2 mm (Gerasimov cited in [29]), that is, sa-sba compared to $g$-ss +2 mm .
(iv) A small and medially directed mastoid process reflects small ears close to the head, and a massive and prominent mastoid process denotes a large and spread ear (9), that is, MLA and MDH compared to ELA, sa-sba, and pra-pa.
(v) The breadth of the ear equals half of its length (9), that is, pra-pa $/$ sa-sba $=0.5$.

TABLE 2—Regression equations for ear prediction.

| Equation No. | Ear Dimension | Formulae | Sex | Population | $n$ | $r^{2}$ | SEE | Data |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| 1 | Length | $(0.22 *$ age $)+55.9$ | Both | U.K. | 206 | UR | UR | (44) |
| 2 | Length | $(0.13 *$ age $)+61.8$ | Both | Japan | 400 | 0.09 | UR | (45) |
| 3 | Length | $(0.12 *$ stature $)+51.2$ | Both | Japan | 400 | 0.04 | UR | (45) |
| 4 | Length/Stature | $(0.00019$ * age) +0.034 | Both | Japan | 400 | 0.36 | UR | (45) |
| 5 | Length | $55.488+0.073 *(\mathrm{FHN}+6)$ | Males | Russia | UR | UR | UR | (49) |
| 6 | Length | $45.650+0.110 *(\mathrm{FHN}+6)$ | Females | Russia | UR | UR | UR | (49) |
| 7 | Length | $\begin{aligned} & (5.89 * \text { sex }) \\ & +(0.21 * \text { age })+52.36 \end{aligned}$ | Both | Mostly U.K. | 2190 | 0.47 | 4.3 | (50) |
| 8 | Length | $\begin{aligned} & (5.06 * \text { sex }) \\ & +(0.15 * \text { age })+55.90 \end{aligned}$ | Both | U.S. White | 1328 | 0.21 | 5.4 | Todd |
| 9 | Width | $\begin{aligned} & (3.04 * \text { sex }) \\ & +(0.05 * \text { age })+33.2 \end{aligned}$ | Both | U.S. White | 1010 | 0.11 | 3.5 | Todd |
| 10 | Length | $\begin{aligned} & (2.13 * \text { sex }) \\ & \quad+(0.16 * \text { age })+54.20 \end{aligned}$ | Both | U.S. Black | 1328 | 0.21 | 5.2 | Todd |
| 11 | Width | $\begin{aligned} & (2.06 * \text { sex }) \\ & +(0.07 * \text { age })+32.9 \end{aligned}$ | Both | U.S. Black | 1010 | 0.19 | 3.1 | Todd |
| 12 | Length | $\begin{aligned} & (4.85 * \text { sex }) \\ & +(0.10 * \text { age })+54.95 \end{aligned}$ | Both | France | 78 | 0.33 | 4.7 | This study |
| 13 | Length | $\begin{aligned} & (\text { sex-0.55)/0.5) } \\ & +(\text { age-41.4)/18.8) } \end{aligned}$ | Both | France | 78 | 0.33 | 6.9 | This study |
| 14 | Width | $\begin{aligned} & (3.20 * \text { sex }) \\ & +(0.05 * \text { age })+33.02 \end{aligned}$ | Both | France | 78 | 0.29 | 3.1 | This study |
| 15 | Length | $\begin{aligned} & (3.98 * \text { sex }) \\ & +(0.12 * \mathrm{FHN})+45.44 \end{aligned}$ | Both | France | 62 | 0.22 | 5.1 | This study |
| 16 | Length | $\begin{gathered} (3.68 * \text { sex })+(0.15 * \text { age }) \\ +(0.14 * \mathrm{FHN})+37.63 \end{gathered}$ | Both | France | 62 | 0.45 | 4.4 | This study |
| 17 | Width | $\begin{gathered} (3.29 * \text { sex })+(0.07 * \text { age }) \\ +(0.06 * \mathrm{FHN})+29.22 \end{gathered}$ | Both | France | 62 | 0.36 | 3.0 | This study |
| 18 | Length | $\begin{aligned} & (4.95 * \text { sex }) \\ & \quad+(0.19 * \text { age })+53.05 \end{aligned}$ | Both | Mixed | 4653 | 0.38 | 5.1 | Combined CT and caliper data of this study, Todd, and Evison and Vorder Bruegge |

FHN, facial height of the skull; age, chronological age in years; sex, dichotomous dummy variable (females $=0$, and males $=1$ ); CT, computed tomography; SEE, standard errors of the estimate; UR, unreported.
${ }^{\dagger}$ To convert standardized predicted ear length to millimeter, employ the formula: Ear Length $(\mathrm{mm})=($ Standardized Predicted Ear Length $* 5.64)+61.94$. This calculation and the unit-weighted regression are based on means and standard deviations observed in the CT scan sample ( $n_{1}$ ).

TABLE 3—Soft tissue landmarks used in this study. All landmarks are defined with the head in the Frankfurt Horizontal.

| Name | Abbreviation | Locality |  |
| :--- | :---: | :--- | :--- |
| Otobasion superius | obs | Bilateral | Pefinition |
| Otobasion inferius | obi | Bilateral | Point of attachment of the helix in the temporal region* |
| Superaurale | sa | Bilateral | Most superior point on the free margin of the auricle* |
| Subaurale | sba | Bilateral | Most inferior point on the free margin of the auricle* |
| Preaurale | pa | Bilateral | Most anterior point of the ear located just in front of the otobasion superius* |
| Postaurale | hx | Milateral | Most posterior point on the free margin of the ear* |
| Helix laterale | se | Midline | Deepest landmark located on the bottom of the nasofrontal angle* |
| Sellion | prn | Midline | Most anterior point of the apex nasi (tip of the nose)* |
| Pronasale | Midline | Apex of the angle at the columella base where the lower |  |
| Subnasale | gidine | border of the nasal septum and the surface of the upper lip meet* |  |
|  |  | Soft tissue analog of hard tissue glabella* |  |
| Soft tissue glabella |  |  |  |

*After (53).
(vi) A strongly developed supramastoid crest corresponds to an upper protrusion of the ear (9), that is, supramastoid crest protrusion compared to EP, EA, and EIA.
(vii) The earlobe is attached if the mastoid process is directed downward and free if the mastoid process points forward (32), that is, earlobe morphology (free or attached) compared to MAA.
(viii) The main axis of the ear is parallel to the angle of the bridge of the nose (27), that is, EA compared to NRA.

This study consequently evaluated a total of eight previously suggested (nonregression) guidelines, two of which have been examined on prior occasions but which were retested here for repeatability purposes (guidelines ii and viii; see [28,33,34,47]). The six other ear prediction guidelines have never been subject to published scientific review.

In order to explore the relationships between the mastoid, ear angles, nose angles, ear heights, and nose heights, a matrix correlation was undertaken in addition to the aforementioned tests.

TABLE 4—Hard tissue landmarks used in this study. All landmarks are defined with the head in the Frankfurt Horizontal.

| Name | Abbreviation | Locality |  |
| :--- | :---: | :--- | :--- |
| Glabella | g | Midline | Definition |
| Nasion | n | Midline | Point located at the intersection of the nasofrontal suture* |
| Rhinion | rhi | Midline | Most anterior point of the internasalis suture* |
| Subspinale | ss | Midline | Apex of curve under the nasal spine* |
| Mastoidale | ms | Bilateral | Most inferior point of the mastoid process* |
| Antero-inferior mastoid | mai | Bilateral | Most antero-inferior point of the mastoid process |
| Menton (or gnathion) | gn | Midline | Lowest median landmark on the lower border of the mandible ${ }^{\dagger}$ |

*After (54).
${ }^{\dagger}$ After (55).


FIG. 2-Ear angles that were measured in this study. See Tables 3 and 5 for definitions (FH = Frankfurt Horizontal; $S=$ Sagittal plane).


FIG. 3-Mandibular ramus and mastoid angles that were measured in this study. See Tables 4 and 5 for definitions ( $F H=$ Frankfurt Horizontal; $C=$ Coronal plane).

Furthermore, the accuracy of the previously mentioned regression equations by Heathcote (44), Asai et al. (45), and Balueva et al. (49) were examined using the CT scan data of this study and, where possible, using the aforementioned samples: (i) 3D face scan data of living subjects collected by Evison and Vorder Bruegge ( $n_{2}=2190$, mean age $=38$ years, $\mathrm{SD}=13$ years, range $=14$ 18 years [50]); (ii) Todd's caliper measurements of White American cadavers from the Hamann Todd Collection that were taken prior to skeletonization $\left(n_{3}=1328\right.$, mean age $=55$ years, $\mathrm{SD}=14$ years, range $=1-96$ years); (iii) Todd's caliper measurements of Black American cadavers from the Hamann Todd Collection that were taken prior to skeletonization ( $n_{4}=1010$, mean age $=41$ years, $\mathrm{SD}=16$ years, range $=1-105$ years); and (iv) a


FIG. 4-Nose angles that were measured in this study. See Tables 3, 4, and 5 for definitions ( $F H=$ Frankfurt Horizontal).


FIG. 5-Supramastoid crest size as morphoscopically evaluated: smooth (left) and strongly developed (right).
group of contemporary living subjects measured using manual anthropometry $\left(n_{5}=47\right.$, mean age $=35$ years, $\mathrm{SD}=9$ years, range $=22-56$ years).

Ear length was calculated from Evison and Vorder Bruegge's data (50) using the "EVB_real_data.csv" file available on the 'Chapter Dataset' CD of Computer-Aided Forensic Facial Comparison after individuals with missing data at landmarks 23 (superaurale) and 25 (subaurale) were removed along with any repeat measurements of the same subjects (i.e., only the first measurement counted). As Evison and Vorder Bruegge only present landmark coordinates, the Euclidean distance between sa and sba was calculated using the Pythagorean formula:


FIG. 6-Earlobe attachment as judged morphoscopically: free (left) and attached (right).

$$
\text { Ear height }=\sqrt{(x a-x b)^{2}+(y a-y b)^{2}+(z a-z b)^{2}}
$$

In the formula, " $x$," " $y$," \& " $z$ " equal the 3D coordinates, $" a "=$ landmark 23 (sa), and " $b "$ = landmark 25 (sba). Note that this measurement may be slightly different from the ear length measurement taken on the CT scan subjects as Evison and Vorder Bruegge do not stipulate that their subjects were positioned in the Frankfurt Horizontal at the time of measurement. In addition, the distances calculated from the 3D coordinate data are not necessarily parallel to the median plane, as they are for the CT data.

With regard to T. W. Todd's unpublished data, the original measurements ( $n=2368$ ) were screened for outliers and missing values. Subsequently, 30 individuals were removed to yield two data sets: one for American Whites $\left(n_{3}=1328\right)$ and one for American Blacks ( $n_{4}=1010$ ). Young individuals (i.e., $1-18$ years of age) were retained in the data sets because (i) ears at birth are close to adult size, that is, represent $75 \%$ of their values at skeletal maturity (56); (ii) linear trends in ear enlargement exist from 10 years of age (56); and (iii) a good spread of young ages was represented in these data sets, hence very young individuals did not represent outliers. Except for being taken manually using sliding calipers, these
data are comparable to the CT scan measurements as they were taken following Martin's Beobachtungsblätter, with the head in the Frankfurt Horizontal (see, e.g., [35]).

Standard errors of the estimate were calculated from residuals of the predicted values obtained from the regression equations using the following equation:

$$
\mathrm{SEE}=\sqrt{\frac{\sum(a-b)^{2}}{n-2}}
$$

where " $a$ " is the true value, " $b$ " is the predicted value, and " $n$ " represents the sample size.

New prediction formulae were generated using stepwise linear regression. To help evaluate the value of the regression formulae, residuals from these equations were compared with those obtained from mean measurements obtained from the CT-derived means using residual plots, SEE (described above), and the $r^{2}$. In addition, crossvalidation was employed using the multiple independent samples described above. To provide comprehensive tests of Balueva et al.'s equations (49), the mean of the in-sample CT scan soft tissue depth at menton (equivalent of Balueva et al.'s gnathion [49]) was calculated and substituted for Balueva et al.'s value of 6 mm (Table 2).

As the CT sample size bordered on the smaller side for stepwise multiple regression using two independent variables ( $n=80$; see [57]), the performance of these regression equations was in addition cross-checked against their improper (unit-weighted) counterparts after recommendations by Cohen et al. (58). These improper regression models were calculated from regression on the standardized values ( $58-60$ ), where +1 was substituted for positive beta weights and -1 was substituted for negative beta weights. This form of regression holds the advantage that large standard errors of the beta weights, which may be encountered in small samples, are avoided $(58,59)$.

Data comparisons were also made against other studies to evaluate consistency of metrics, such as for ear length. Statistical calculations were undertaken using STATISTICA ${ }^{\circledR}$ (v. 7.1; Statsoft ${ }^{\circledR}$, Tulsa, OK), SPSS ${ }^{\circledR}$ (v. 11.0; IBM ${ }^{\circledR}$, Somers, NY), and Microsoft ${ }^{\circledR}$ Excel $^{\circledR}$ (2007; Microsoft ${ }^{\circledR}$, Redmond, WA).

TABLE 5-Measurements taken in this study.

| Name | Abbreviation | Locality | Zone | Definition |
| :---: | :---: | :---: | :---: | :---: |
| Ear height | sa-sba | Bilateral | ST | Distance from superaurale to subaurale |
| Ear width | pra-pa | Bilateral | ST | Distance from preaurale to postaurale |
| Ear insertion height | obs-obi | Bilateral | ST | Distance from otobasion superius to otobasion inferius |
| Ear angle | EA | Bilateral | ST | Angle between the main axis of the ear in a lateral view and the FH |
| Ear insertion angle | EIA | Bilateral | ST | Angle between the axis of the ear insertion (otobasions) in a lateral view and the FH |
| Ear protrusion | EP | Bilateral | ST | Angle between the main axis of the ear in a superior view and the sagittal plane |
| Nasal dorsum length | se-prn | Midline | ST | Distance from sellion to pronasale |
| Nasal root angle | NRA | Midline | ST | Angle between a line formed by the sellion and the pronasale, and the FH |
| Soft nose height 1 | se-sn | Midline | ST | Distance from sellion to subnasale |
| Soft nose height 2 | g'-sn | Midline | ST | Distance from soft glabella to subnasale |
| Mastoid height | MDH | Bilateral | HT | Height of the mastoid from the mastoidale and an orthogonally projected point on the FH |
| Mastoid anterior angle | MAA | Bilateral | HT | Angle between the main axis of the mastoid process in a lateral view and the FH |
| Mastoid lateral angle | MLA | Bilateral | HT | Angle between the main axis of the mastoid process in an inferior view and the coronal plane |
| Mandible ramus angle | MRA | Bilateral | HT | Angle between the main axis of the posterior part of the mandibular ascending ramus and the FH |
| Hard nose angle | HNA | Midline | HT | Angle between a line formed by the nasion and the rhinion, and the FH |
| Hard nose height 1 | n -ss | Midline | HT | Distance from nasion to subspinale |
| Hard nose height 2 | g-ss | Midline | HT | Distance from glabella to subspinale |
| Facial height | FHN | Midline | HT | Distance from nasion to menton |

ST, soft tissue; HT, hard tissue; FH, Frankfurt Horizontal; FHN, facial height of the skull.

TABLE 6—Descriptive statistics, correlations with age, and statistical significance tests for asymmetry and sexual dimorphism.

| Measurement | $n$ | Mean | Min | Max | SD | Asymmetry ( $p$-Value) |  |  | Sex |  |  |  |  |  | Age (r) |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | M \& F | M | F | Mean M | Mean F | SD M | SD F | $t$ | $p$ |  |
| Angles (degrees) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| MAA [R] | 78 | 63.6 | 30.7 | 85.2 | 12.0 | $<0.01$ | 0.05 | 0.03 | 62.1 | 65.4 | 12.3 | 11.5 | -1.21 | 0.23 | -0.19 |
| MAA [L] | 78 | 60.2 | 29.9 | 84.5 | 10.4 |  |  |  | 59.1 | 61.6 | 11.0 | 9.6 | -1.05 | 0.29 | 0.01 |
| MLA [R] | 78 | 109.6 | 79.7 | 130.7 | 9.7 | 0.01 | <0.01 | 0.68 | 107.4 | 112.4 | 9.8 | 8.9 | -2.34 | 0.03 | -0.03 |
| MLA [L] | 78 | 112.8 | 85.2 | 140.1 | 11.0 |  |  |  | 113.7 | 111.7 | 11.8 | 9.9 | 0.81 | 0.42 | 0.12 |
| MRA [R] | 78 | 85.8 | 72.8 | 90.0 | 3.9 | 0.73 | 0.83 | 0.43 | 85.4 | 86.3 | 4.0 | 3.7 | -1.03 | 0.23 | 0.00 |
| MRA [L] | 78 | 85.7 | 73.2 | 90.0 | 4.4 |  |  |  | 85.4 | 86.0 | 4.5 | 4.3 | -0.52 | 0.38 | 0.00 |
| EP [R] | 78 | 48.5 | 17.5 | 74.6 | 11.8 | 0.76 | 0.85 | 0.56 | 47.4 | 49.7 | 10.9 | 12.9 | -0.85 | 0.40 | 0.33 |
| EP [L] | 78 | 48.8 | 20.0 | 89.0 | 11.4 |  |  |  | 47.2 | 50.7 | 10.9 | 11.8 | -1.37 | 0.17 | 0.21 |
| EA [R] | 78 | 77.2 | 59.6 | 89.6 | 6.3 | <0.01 | 0.33 | <0.01 | 76.9 | 77.6 | 6.1 | 6.6 | -0.47 | 0.64 | 0.02 |
| EA [L] | 78 | 75.7 | 60.7 | 89.3 | 6.1 |  |  |  | 76.0 | 75.2 | 6.5 | 5.8 | 0.59 | 0.55 | 0.19 |
| EIA [R] | 78 | 81.0 | 68.2 | 95.0 | 5.9 | 0.71 | 0.84 | 0.75 | 80.7 | 81.3 | 5.6 | 6.2 | -0.42 | 0.68 | 0.41 |
| EIA [L] | 78 | 81.3 | 69.8 | 97.6 | 6.0 |  |  |  | 80.9 | 81.7 | 5.7 | 6.4 | -0.58 | 0.56 | 0.40 |
| NRA | 77 | 58.2 | 48.6 | 72.5 | 3.9 | - | - | - | 58.3 | 58.1 | 4.2 | 3.7 | 0.18 | 0.86 | 0.08 |
| HNA | 77 | 53.3 | 32.5 | 75.2 | 7.0 | - | - | - | 52.6 | 54.2 | 7.1 | 6.8 | -1.05 | 0.30 | 0.13 |
| MDH [R] | 78 | 30.0 | 20.2 | 38.0 | 4.1 | 0.04 | 0.36 | 0.04 | 31.8 | 27.7 | 3.3 | 3.9 | 4.96 | <0.01 | 0.02 |
| MDH [L] | 78 | 29.5 | 18.7 | 40.5 | 4.0 |  |  |  | 31.5 | 27.1 | 3.2 | 3.4 | 5.87 | <0.01 | 0.00 |
| Linear distances (mm) |  |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| se-prn | 72 | 48.6 | 37.2 | 58.7 | 4.9 | - | - | - | 50.2 | 46.6 | 4.7 | 4.5 | 3.32 | 0.01 | 0.10 |
| se-sn | 71 | 54.7 | 44.9 | 66.7 | 4.8 | - | - | - | 56.1 | 53.1 | 4.5 | 4.7 | 2.64 | 0.01 | 0.05 |
| g'-sn | 73 | 67.1 | 56.5 | 79.1 | 4.6 | - | - | - | 68.0 | 66.2 | 4.6 | 4.5 | 1.67 | 0.10 | 0.11 |
| n -ss | 77 | 56.1 | 48.8 | 65.8 | 3.4 | - | - | - | 57.6 | 54.4 | 3.4 | 2.6 | 4.59 | <0.01 | -0.06 |
| g -ss | 77 | 66.4 | 59.8 | 75.7 | 3.7 | - | - | - | 66.7 | 65.9 | 3.7 | 3.7 | 0.97 | 0.33 | -0.14 |
| obs-obi [R] | 78 | 54.8 | 41.7 | 69.7 | 5.5 | 0.27 | 0.90 | 0.09 | 57.2 | 51.7 | 5.6 | 3.6 | 5.05 | <0.01 | 0.58 |
| obs-obi [L] | 78 | 55.1 | 44.1 | 71.2 | 5.5 |  |  |  | 57.2 | 52.4 | 6.1 | 3.2 | 4.19 | <0.01 | 0.62 |
| sa-sba [R] | 78 | 61.7 | 49.5 | 74.3 | 4.9 | 0.34 | 0.31 | 0.84 | 63.8 | 59.0 | 4.8 | 3.7 | 4.92 | <0.01 | 0.47 |
| sa-sba [L] | 78 | 61.9 | 48.2 | 80.1 | 5.6 |  |  |  | 64.3 | 59.1 | 5.5 | 4.3 | 4.52 | <0.01 | 0.50 |
| pra-pa [R] | 78 | 36.6 | 29.7 | 43.9 | 3.2 | 0.40 | 0.33 | 0.73 | 37.9 | 34.9 | 2.8 | 2.9 | 4.67 | <0.01 | 0.33 |
| pra-pa [L] | 78 | 37.0 | 30.9 | 45.3 | 3.6 |  |  |  | 38.5 | 35.2 | 3.2 | 3.3 | 4.58 | <0.01 | 0.32 |
| FHN | 62 | 116.9 | 97.2 | 133.2 | 8.5 | - | - | - | 120.9 | 111.4 | 7.2 | 6.9 | 5.21 | 0.04 | -0.02 |

[R], right; [L], left; M, males; F, females.
Bold type $=$ statistically significant at $p<0.05$.

## Results

Shapiro-Wilk's test indicated that all metric measurements of the CT scans (Table 5) followed normal distributions ( $p<0.05$ ), except for the mandible angle. Wilcoxon test, Mann-Whitney $U$-test, and Spearman correlations were, therefore, used where this measurement was considered. All other variables were subject to parametric statistics.

Table 6 presents the descriptive statistics of the measurements from the CT scans, statistical significance test results, and correlations concerning asymmetry, sexual dimorphism, and age. Of note, a large and statistically significant difference existed between males and females for ear length and width (male mean, left side $=64.3 \mathrm{~mm}, \quad \mathrm{SD}=5.5, n=43$; female mean, left side $=$ $59.1 \mathrm{~mm}, \mathrm{SD}=4.3 \mathrm{~mm}, n=35 ; p<0.01$ ). Age was also positively correlated with these ear dimensions (left side, sexes combined, $r=0.50$ and 0.32 , respectively; see also Fig. 7), but variability in measurements between individuals was large (one standard deviation represented c. $10 \%$ of the total ear length measurement). From this point forward, only the data for the left side of the face will be addressed.

The CT ear length measurements corresponded closely to those calculated from Evison and Vorder Bruegge's data (50), Todd's measurements on American Blacks, and the measurements on contemporary living subjects (compare to Table 7). The mean CT ear lengths were, however, c. 5 mm less than those calculated from the Todd data for U.S. White cadavers (Table 7).

In regards to the eight previously published ear prediction rules that were examined in this study, the following results were observed using the CT scan data:

Rule (i): The posterior jaw line (MRA) was not parallel to the orientation of the ear (EA) and these two variables were not correlated (Table 8).
Rule (ii): The height of the nose (se-sn) underestimated the ear height (sa-sba) and the height from $g$ ' to sn overestimated the height of the ear by 5.5 mm (Table 6). Positive correlations existed between the height of the nose and the ear; however, they were weak ( $r=0.30$; Table 8 ).
Rule (iii): The addition of a further 2 mm to the g 'sn measurement, as Gerasimov recommended, produced additional error in comparison to his recommendations tested at rule (ii).
Rule (iv): The medio-lateral orientation of the mastoid process (MLA) did not correlate with the size and orientation of the outer ear; however, there was a small correlation between the MDH and the breadth of the ear $(r=0.41$; Table 9).
Rule (v): The breadth of the ear (pra-pa) was not equal to half its length; however, length and width were correlated $(r=0.60)$. The mean ratio of the breadth of the ear to its height was 0.59 , a number which compares favorably to that reported in other contemporary studies: for example, $0.60-0.61$ (39), 0.59-0.62 (41), $0.55-0.57$ (43), $0.56-0.58$ (34), and $0.56-0.57$ (36).

Rule (vi): Development of the supramastoid crest was not related to the upper protrusion of the ear (EP); however, the protrusion of the crest is significantly linked to the ear height (sa-sba and obs-obi) and


FIG. 7-Scatterplots of the ear length and width (left side) with age.

TABLE 7-Sex-specific mean ear dimensions in the four samples evaluated in this study.

| Investigator | Measurement Method | Sex | Population | Age (Years) | $n$ | Ear Length |  |  |  |  | Ear Width |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  |  | Mean |  | Min | Max | $r$ (age) | Mean | SD | Min | Max | (age) |
| Evison and Vorder Bruegge (50) | 3D Face Scanner | Males | Mostly U.K. | 14-81 | 1401 | 66.4 |  | 49.0 | 83.1 | 0.6 | - | - | - | - | - |
| Todd | Anthropometry |  | American Whites | 1-96 | 1160 | 69.4 |  | 40.0 | 91.0 | 0.4 | 39.0 | 3.5 | 23.0 | 51.0 | 0.2 |
| Todd | Anthropometry |  | American Blacks | 1-105 | 758 | 62.9 |  | 40.0 | 83.0 | 0.4 | 38.0 | 3.4 | 24.0 | 49.0 | 0.4 |
| This study | Anthropometry |  | Mixed | 24-52 | 22 | 65.1 | 4.9 | 53.0 | 77.0 | 0.6 | 36.1 | 2.7 | 32.0 | 42.5 | 0.3 |
| Evison and Vorder Bruegge (50) | 3D Face Scanner | Females | Mostly U.K. | 15-76 | 789 | 59.8 | 4.5 | 44.2 | 76.1 | 0.4 | - | - | - | - | - |
| Todd | Anthropometry |  | American Whites | 1-93 | 168 | 64.4 |  | 40.0 | 77.0 | 0.5 | 36.0 | 3.8 | 24.0 | 47.0 | 0.4 |
| Todd | Anthropometry |  | American Blacks | 1-89 | 252 | 60.4 | 5.9 | 45.0 | 75.0 | 0.5 | 35.8 | 3.1 | 26.0 | 45.0 | 0.4 |
| This study | Anthropometry |  | Mixed | 22-56 | 25 | 61.3 | 4.1 | 55.0 | 69.0 | 0.3 | 32.5 | 2.4 | 29.0 | 37.0 | 0.2 |

TABLE 8—Ear axis and ear length relationships with ramus axis and nose length measurements (t-tests and Pearson's correlations).

| a | b | Mean a | Mean b | SD a | SD b | $t$ | $p$ | $r$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EA | MRA | 75.7 | 85.7 | 6.1 | 4.4 | -11.71 | $<0.01$ | 0.12 |
| sa-sba | se-sn | 61.9 | 54.7 | 5.6 | 4.8 | 8.34 | $<0.01$ | 0.25 |
| sa-sba | n-ss | 61.9 | 56.1 | 5.6 | 3.4 | 7.76 | $<0.01$ | 0.29 |
| sa-sba | g-ss | 61.9 | 66.4 | 5.6 | 3.7 | -5.76 | $<0.01$ | 0.01 |

TABLE 9—Correlation matrix (Pearson's r) between the mastoid process and the outer ear.

|  | EP [L] | sa-sba [L] | pra-pa [L] |
| :--- | :---: | :---: | :---: |
| MLA [L] | -0.06 | 0.04 | -0.07 |
| MDH [L] | -0.12 | 0.26 | 0.41 |

[L], left side.
orientation at its insertion (EIA). That is, individuals with protruded supramastoid crests had on average larger sa-sba, obs-obi, and EIA measurements (Table 10). In this study's sample, a strongly developed supramastoid crest was observed in 33 participants $(42 \%)$ and a smooth supramastoid crest was observed in the other 45 subjects ( $58 \%$ ). In general, no right/left asymmetry was noted.
Rule (vii): The lobe morphology (free or attached) was not related to the mastoid form or other ear measurements (Table 11). Free lobes were observed in 55 individuals ( $71 \%$ ) and the other 23 subjects ( $29 \%$ ) had attached lobes. No overarching right/left asymmetry was observed.
Rule (viii): The main axis of the ear (EA) was not parallel to the angle of the soft (NRA) or hard nose (HNA),

TABLE 10—Relationship between the protrusion of the supramastoid crest and the outer ear.

|  | Supramastoid Crest Protrusion |  |  |  |  |  |
| :--- | :--- | ---: | :--- | ---: | ---: | ---: |
|  | Protruded <br> $(n=33)$ |  | Smooth $(n=45)$ |  |  |  |
|  | Mean | SD |  | Mean | SD | $t$-Test $(p)$ |
| EP [L] | 45.9 | 11.1 |  | 50.9 | 11.2 | 0.05 |
| EA [L] | 74.3 | 6.3 | 76.7 | 5.8 | 0.08 |  |
| EIA [L] | 80.7 | 5.2 | 81.7 | 6.6 | 0.47 |  |
| obs-obi [L] | 57.2 | 6.0 | 53.5 | 4.6 | $<\mathbf{0 . 0 1}$ |  |
| sa-sba [L] | 64.0 | 6.1 | 60.4 | 4.8 | $<\mathbf{0 . 0 1}$ |  |
| pra-pa [L] | 38.0 | 3.3 | 36.3 | 3.7 | $\mathbf{0 . 0 3}$ |  |

[^1]and the ear and nose orientation did not correlate (Table 12). The mean difference between the ear and soft tissue nose angle in this study was $18^{\circ}$ slightly higher than that of other studies ( $33,34,47,48$ ) -and probably as a result of defining the soft tissue nose angle from sellion to pronasale as opposed to directly running along the nasal dorsum as in other studies. The difference between the ear angulation and the angle of the hard tissue bridge of the nose was larger than the difference between the ear angulation and the angle of the soft tissue nose (mean $=23^{\circ}$ ).

A cross-table analysis of the supramastoid crest and the lobe morphology (Table 13) also indicated a statistically significant relationship between the crest protrusion and the lobe morphology (Pearson's $\chi^{2}<0.05$ ). When the crest is protruded, the lobe tends to be free. However, the inverse is not true: the attached lobe is not linked with a smooth supramastoid crest. A similar analysis showed no difference between males and females for the lobe attachment ( $p=0.87$, contingency coefficient $=0.01$ ), but a strong relationship was detected between sex and supramastoid crest protrusion ( $p<0.01$, contingency coefficient $=0.46$ ). In all, $86 \%$ of the females had a smooth crest versus $65 \%$ of the males.
In this sample, the orientation of the ear (EP and EA) did not correlate with the mastoid process or the bony regions of the nose (see results of matrix correlation, Table 14). The MDH correlated with the ear breadth (pra-ba), and the height of the piriform aperture ( n -ss) only weakly correlated with the ear dimensions (sa-sba, pra-pa).

On the basis of the relationships between ear dimensions and sex, age, and facial height of the skull (FHN) (see above and the example scatter plots provided in Fig. 7), 12 new prediction equations were generated using stepwise regression (Table 2). In all cases, sex was transformed to a dichotomous dummy variable before entering the regression analysis (females $=0$, and males $=1$ ). Because mean soft tissue depths at menton in the CT sample differed by $\leq 3.5 \mathrm{~mm}$ from Balueva et al.'s (49) recommendation of 6 mm males $=9.5 \mathrm{~mm}, n=36$; females $=8.1 \mathrm{~mm}$, $n=26$ ), revised versions of the Balueva et al. equations using actual soft tissue depth values found in the sample were not pursued further ( $3.5 \mathrm{~mm} * 0.073=0.3 \mathrm{~mm}$ error).

Of the 18 regression equations evaluated in total in this study, those that disregarded sex and concerned only age produced the highest SEE (Tables 2 and 15). The ear length equation generated from Evison and Vorder Bruegge's data (no. 7) performed well on the CT scan data, but SEE was slightly lower when the equation generated from the CT scan data of this study (no. 12) was applied to Evison and Vorder Bruegge's data set. Equation no. 7 performed better than equation no. 12 on the U.S. White cadaver data and that derived from caliper measurements, but equation no. 7 performed much worse on

TABLE 11—Relationship between the earlobe morphology and the ear/mastoid features.

|  | Lobe Morphology |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Attached$(n=23)$ |  | Free ( $n=55$ ) |  | $t$-Test (p) |
|  | Mean | SD | Mean | SD |  |
| MAA [L] | 60.4 | 11.0 | 60.1 | 10.2 | 0.90 |
| MLA [L] | 113.9 | 9.6 | 112.4 | 11.6 | 0.59 |
| MDH [L] | 29.7 | 4.6 | 29.5 | 3.7 | 0.85 |
| EP [L] | 48.9 | 12.7 | 48.7 | 11.0 | 0.96 |
| EA [L] | 77.6 | 6.8 | 74.9 | 5.7 | 0.08 |
| EIA [L] | 82.5 | 7.1 | 80.8 | 5.5 | 0.26 |
| obs-obi [L] | 54.5 | 6.1 | 55.3 | 5.3 | 0.53 |
| sa-sba [L] | 61.8 | 5.5 | 62.0 | 5.8 | 0.86 |
| pra-pa [L] | 37.1 | 3.6 | 37.0 | 3.7 | 0.96 |

[L], left side.

TABLE 12—Relationships between the ear orientation and the angle of the nose ( t -test and Pearson's correlations).

| a | b | Mean a | Mean b | SD a | SD b | t | $p$ | $r$ |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| EA [L] | NRA | 75.7 | 58.2 | 6.1 | 3.9 | 21.13 | $<0.01$ | 0.02 |
| EA [L] | HNA | 75.7 | 53.3 | 6.1 | 7.0 | 21.22 | $<0.01$ | 0.08 |

[^2]TABLE 13—Cross-table analysis of the supramastoid crest protrusion and the earlobe morphology.

|  | Supramastoid Crest Protrusion |  |
| :--- | :---: | ---: |
|  | Developed <br>  <br> Lobe | Smooth <br> $n(\%)$ |
| Attached | $5(15)$ | $18(40)$ |
| Free | $28(85)$ | $27(60)$ |
| Total | $33(100)$ | $45(100)$ |

the U.S. Black cadaver sample (Table 15). Ear length equations derived on the cadaver samples (no. 8 and 10) performed well on other samples, but generally had higher SEE on the in- and out-ofsample groups than equation no. 7 or 12 (Table 15).

Ear length prediction equations of Balueva et al. (no. 5 and 6) produced similar SEE on the CT scan data as other equations reported above, but $r^{2}$ values were extremely low ( $<0.08$; Table 15). In addition, the in-sample CT scan data indicated that FHN correlated less strongly with ear length, than did age ( $r^{2}=0.37$ and 0.48 , respectively), and regression equations based on sex and FHN ( $\mathrm{SEE}=5.1, r^{2}=0.22$ ) performed worse than that based on both sex and age ( $\mathrm{SEE}=4.7, r^{2}=0.33$; Table 2). Although $r^{2}$ values increased, with decreasing SEE, for FHN regression equations that included age (Table 2; compare equations no. 16 and 12), overfitting is possible because the ratio between sample size and the number of independent variables is further reduced (from $40-1$ to $26-1$, see [57-59]). For these reasons, and more importantly because mandible position must be estimated in edentulous skulls yielding additional error, any of the regression equations that include FHN as an independent variable are not recommended.

The unit-weighted regression equation for ear length (no. 13) did not perform more accurately than the proper regression equation (no. 12), even in contrast to cross-validated results, and therefore will not be further discussed here. Sex-specific mean ear lengths, derived from the CT scans, generally predicted ear length

TABLE 14-Correlation matrix (Pearson's r) between the outer ear and the bony features of the ear and nose.

|  | EP [L] | EA [L] | sa-sba [L] | pra-pa [L] | obi-obs [L] |
| :--- | :---: | :---: | :---: | :---: | :---: |
| MAA [L] | 0.07 | 0.12 | -0.01 | -0.02 | 0.00 |
| MLA [L] | -0.06 | 0.17 | 0.05 | -0.07 | 0.03 |
| MRA [L] | 0.23 | 0.15 | -0.09 | -0.12 | 0.06 |
| HNA | 0.19 | 0.08 | -0.02 | -0.03 | -0.01 |
| MDH [L] | -0.13 | -0.01 | 0.23 | $\mathbf{0 . 4 0}$ | 0.10 |
| n-ss | 0.01 | 0.13 | $\mathbf{0 . 3 1}$ | $\mathbf{0 . 3 6}$ | 0.14 |
| g-ss | 0.18 | 0.04 | 0.07 | 0.12 | -0.01 |

[^3]TABLE 15-Regression equation performances on different samples.

| Ear Dimension | Equation No. (see Table 2) |  |  |  | Out-of-Sample Data |  |  |  |  |  |  |  |  |  |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | In-S | ample | Data | G/S <br> CT Scans Living France |  |  | E/VB <br> 3D Face Scans Living White British |  |  | Todd <br> Calipers <br> Cadavers <br> U.S. White |  |  | Todd Calipers Cadavers U.S. Black |  |  | G/S <br> Calipers <br> Living <br> Mixed |  |  |
|  |  | SEE | $r^{2}$ | $n$ | SEE | $r^{2}$ | $n_{1}$ | SEE | $r^{2}$ | $n_{2}$ | SEE | $r^{2}$ | $n_{3}$ | SEE | $r^{2}$ | $n_{4}$ | SEE | $r^{2}$ | $n_{5}$ |
| Length | 1 | UR | UR | 206 | 6.4 | 0.15 | 78 | 5.2 | 0.25 | 2190 | 5.8 | 0.13 | 1328 | 6.0 | 0.19 | 1010 | 4.4 | 0.21 | 47 |
|  | 2 | UR | 0.09 | 400 | 7.5 | 0.15 | 78 | 6.0 | 0.25 | 2190 | 5.7 | 0.13 | 1328 | 7.2 | 0.19 | 1010 | 5.4 | 0.21 | 47 |
|  | 5 | UR | UR | UR | 5.4 | 0.08 | 36 | - | - | - | - | - | - | - | - | - | - | - | - |
|  | 6 | UR | U | UR | 5.0 | 0.00 | 26 | - | - | - | - | - | - | - | - | - | - | - | - |
|  | 7 | 4.3 | 0.47 | 2190 | 5.6 | 0.31 | 78 | - | - | - | 6.2 | 0.21 | 1328 | 8.0 | 0.20 | 1010 | 4.2 | 0.32 | 47 |
|  | 8 | 5.4 | 0.21 | 1328 | 5.6 | 0.32 | 78 | 4.5 | 0.48 | 2190 | - | - | - | - | - | - | 4.1 | 0.31 | 47 |
|  | 10 | 5.2 | 0.21 | 1328 | 5.0 | 0.25 | 78 | 5.0 | 0.42 | 2190 | - | - | - | - | - | - | 4.5 | 0.35 | 47 |
|  | 12 | 4.7 | 0.33 | 78 | - | - | - | 5.1 | 0.45 | 2190 | 6.8 | 0.20 | 1328 | 5.4 | 0.16 | 1010 | 4.8 | 0.27 | 47 |
|  | 13 | 6.9 | 0.33 | 78 | - | - | - | 5.6 | 0.46 | 2190 | 6.9 | 0.20 | 1328 | 6.3 | 0.18 | 1010 | 6.7 | 0.30 | 47 |
| Width | 9 | 3.5 | 0.11 | 1010 | 3.1 | 0.33 | 78 | - | - | - | - | - | - | - | - | - | 3.3 | 0.40 | 47 |
|  | 11 | 3.1 | 0.19 | 1010 | 4.2 | 0.25 | 78 | - | - | - | - | - | - | - | - | - | 2.7 | 0.36 | 47 |
|  | 14 | 3.1 | 0.29 | 78 | - | - | - | - | - | - | 3.4 | 0.11 | 1328 | 3.2 | 0.17 | 1010 | 3.2 | 0.40 | 47 |

E/VB, Evison and Vorder Bruegge (50); G/S, Guyomarc'h and Stephan (this study); UR, unreported; CT, computed tomography; SEE, standard errors of the estimate.

G/S CT Mean
G/S CT Scans












Equation \#12

FIG. 8-Residual plots of estimated ear lengths, by age, using the $G / S$ computed tomography $(C T)$ mean and equation no. 12. The slopes represent linear trend lines. For equation no. 12 in the G/S CT sample, this trend line falls on precisely on the x-axis.
well (Fig. 8); however, regression equation no. 12 (and no. 7) did much better at predicting those individuals at the extremes of the age ranges ( $2-10 \mathrm{~mm}$ less error; Fig. 8). It should be noted, however, that this equation generally tended to underestimate true ear length values across the out-of-sample groups, except for extremely young individuals, whose ear length was overestimated (see Fig. 8). A regression equation for ear length using all the raw data examined in this study is presented in Table 2 (equation no. 18, $n=4653$ ), but unfortunately cross-validation is not possible owing to a lack of independent data for tests. Nevertheless, robust results are expected, given the large sample size upon which this equation is derived.

With regard to ear width, the CT scan-derived equation (no. 14) performed almost as well on Todd's data and the caliper measured contemporary sample as the in-sample formulae (Table 15). However, residuals between predicted and actual ear width values indicated no value of the CT-derived regression equation above the arithmetic mean as a predictor (Fig. 9). Consequently, mean ear widths are recommended for ear width prediction above the regression equations.

## Discussion

Tests of previously published face prediction rules have been conducted for many of the major facial regions (eyes, nose, and mouth), except for the ear. This study fills that gap and its results are consistent with a general pattern observed in other studies that traditionally recommended face prediction rules are not well supported by the scientific data ( $10-18,61$ ).

The inaccuracies of previously published soft tissue prediction guidelines are not ambiguous or limited to few specific facial regions. Similar findings have been found on numerous occasions
and often by independent teams of investigators, see, for example, results on eyeball position (13-15,61,62). In this context, the poor performance of overarching face prediction methods (see, e.g., [26,63-65]) is not surprising and mandates improvements to facilitate accurate face prediction.

This study makes some progress toward achieving this goal with specific regard to the ear. First, faces of non-Asian individuals should be constructed with free earlobes because (i) valid skeletal indicators for attached earlobes do not currently exist; (ii) supramastoid crests are associated with free earlobes (Table 13); and (iii) free earlobes tend to be more common in non-Asian groups (Tables 13 and 16). Second, orientations of the ear should follow quantified published means, either of this study (Table 6) or those published elsewhere (see, e.g., [33,34,47,48]). Third, ear length should be predicted using the proper regression equation derived from CT scans in this study (no. 12) because, so far as the data currently suggest, it provides the greatest generality that has been verified by cross-validation (Table 15 and Fig. 8). Fourth, for ear width prediction, the mean ear width should be used as it provides a simple and accurate estimator whose performance is not surpassed by the regression equations so far examined (Fig. 9). At this stage, all other characteristics of the ear can be derived from the skull only by speculation.

Although the tested guidelines presented in this study will assist facial approximation, they should not be interpreted to mean that prediction of the ear is an accurate endeavor. The magnitude of errors associated with the aforementioned four tested ear prediction guidelines is large and, in conjunction with the paucity of tested methods for ear prediction in general, underscores the inexact nature of the ear prediction process. Error in age prediction from the skeleton will also inflate errors associated with ear length prediction using the aforementioned regression equation. These inaccuracies,


FIG. 9-Residual plots of estimated ear widths, by age, using the G/S computed tomography (CT) mean and equation no. 14. The slopes represent linear trend lines.
errors, and gaps in knowledge unambiguously highlight the approximate nature of the face prediction process and the sense to abandoning the term "facial reconstruction" in favor of "facial approximation."
The limited interrelationships and physical connection between the soft tissues of the ear and the temporal bone, as measured in this study, will likely hamper future attempts to improve ear prediction guidelines. Nevertheless, the large amount of variance unexplained by sex and age justifies future searches for variables that offer improvements. Moreover, landmark- and outline-based morphometric studies should be considered, as should ear relationships to other more-distant parts of the skull.

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TABLE 16-Morphotype frequencies of the earlobe in the human population (\%).

| Population | Attached | Free | Intermediate | Sample | Subjects | $n$ | Study |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Caucasoids | 29 | 71 | - | French | Unrelated | 78 | This study |
|  | 35 | 65 | - | Indians (Indian subcontinent) | Not specified | 111 | (67) |
|  | 14 | 64 | 22 | Indians (Indian subcontinent) | Not specified | 97 | (68) |
|  | 17 | 56 | 27 | Indians (Indian subcontinent) | Not specified | 103 | (68) |
|  | 13 | 50 | 37 | Indians (Indian subcontinent) | Not specified | 54 | (68) |
|  | 26 | 52 | 22 | Indians (Indian subcontinent) | Not specified | 54 | (68) |
|  | 25 | 49 | 25 | Indians (Indian subcontinent) | Not specified | 102 | (68) |
|  | 24* | 77 | - | Indians (Indian subcontinent) | Unrelated | 107 | (69) |
|  | 31 | 69 | - | Indians (Indian subcontinent) | Unrelated | 80 | (69) |
|  | 23 | 34 | 43 | Indians (Indian subcontinent) | Not specified | 1288 | (70) |
|  | 16 | 50 | 35 | Indians (Indian subcontinent) | Not specified | 210 | (70) |
|  | 24 | 76 | - | Indians (Indian subcontinent) | Not Specified | 100 | (71) |
|  | 34 | 66 | - | Indians (Indian subcontinent) | Not Specified | 119 | (72) |
|  | 25 | 75 | - | Indians (Indian subcontinent) | Not Specified | 183 | (72) |
|  | 41 | 59 | - | North Americans | Not specified | 380 | Glass et al. (1952) cited in (69) |
|  | 25 | 75 | - | North Americans | Not specified | 241 | Glass et al. (1952) cited in (69) |
|  | 84 | 16 | - | North Americans | Families | 248 | (73) |
|  | 18 | 54 | 28 | North Americans | Not specified | 381 | (74) |
|  | 16 | 84 | - | North Brazilians | Not specified | - | Saldanha (66) cited in (69) |
|  | 23 | 77 | - | Scottish | Not specified | 500 | (75) |
|  | 35 | 65 | - | Swedish | Families | 247 | (76) |
|  | 33 | 67 | - | Yugoslavs | Families | - | Berberovic and Hadziselimovic (1972) cited in (77) |
| Negroids | 25 | 75 | - | Nigerian | Unrelated | 1600 | (77) |
|  | 32 | 48 | 32 | North Americans | Not specified | 242 | (74) |
| Mongoloids | 62 | 38 | - | Ahom | Unrelated | 100 | (78) |
|  | 29 | 71 | - | Ahom | Not specified | 330 | Das (1975) cited in (78) |
|  | 64 | 36 | - | Chinese | Not specified | 79 | (67) |
|  | $75^{\dagger}$ | 25 | - | Fijians | Not Specified | 813 | (46) |
|  | 65 | 35 | - | Filipino | Not specified | 49 | (67) |
|  | 46 | 54 | - | Garo | Not specified | 200 | Das (1967) cited in (78) |
|  | 67 | 33 | - | Japanese | Not specified | 70 | (67) |
|  | 34 | 66 | - | Kachari | Not specified | 100 | Das (1967) cited in (78) |
|  | 23 | 77 | - | Kalita | Not specified | 120 | Das (1964) cited in (78) |
|  | 49* | 51 | - | Newars | Not Specified | 169 | (80) |
|  | 45 | 18 | 37 | Nicobar Islands-Car | Not specified | 341 | (68) |
|  | 55 | 14 | 31 | Nicobar Islands-Chowrite | Not specified | 111 | (68) |
|  | 27 | 37 | 37 | Nicobar Islands-Terressan | Not specified | 146 | (68) |
|  | 65 | 10 | 25 | Nicobar Islands-Southern | Not specified | 119 | (68) |
|  | 64 | 36 | - | Papua New Guinean | Families | 399 | (67) |
|  | 23 | 77 | - | Rabha | Not specified | 300 | Das (1967) cited in (78) |
|  | 26 | 74 | - | Rajbansi | Not specified | 100 | Das (1967) cited in (78) |
|  | 63 | 37 | - | Sema Naga | Not specified | 100 | Phukan and Begum (1976) cited in (78) |
|  | 21 | 79 | - | Suri | Not specified | 100 | Das and Deka (1960) cited in (78) |
|  | 57 | 22 | 20 | Tibetan | Not specified | 250 | (79) |
| Australoids | 22 | 78 | - | Australian Aborigines | Not specified | 41 | (67) |
|  | 50 | 18 | 32 | Onge | Not specified | 80 | (81) |

[^4]${ }^{\dagger}$ Includes $10 \%$ reported as "soldered" in the original study.
draft of this manuscript that inspired us to undertake cross-validation using independent samples.

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[^1]:    [L], left side.
    Bold type $=$ statistically significant at $p<0.05$.

[^2]:    ## [L], left side.

[^3]:    [L], left side.
    Bold type $=$ statistically significant at $p<0.05$.

[^4]:    *Includes 5/6\% reported as "lobe-less" in the original study.

